Small-Area Life Expectancy: A Comparison of Methods, Relationship to Neighborhood Sociodemographic Factors and Outlier Analysis

Life expectancy (LE) has become increasingly popular as a tool for demonstrating health disparities across countries, US states, and small geographic areas where area-level characteristics may, in part, account for variability in health outcomes. LE varies substantially across U.S. communities and even small areas such as neighborhoods [1, 2]. However, numerous methodological challenges arise when life table methods, which are well suited to larger populations, are applied to smaller populations [3-5]. Additionally, localized, place-based characteristics play an important role in LE and other measures of population health. To date, only a few studies have used a fine geographic scale to identify local assets that might enhance health outcomes [6-8]. Recently, there has been increased interest by policymakers and health officials to improve and refine methods to assess those small-area factors that influence population and neighborhood health. Likewise, there is also an increased need to understand outlier communities—areas that have unusually high or low LE or other health measures that would not be predicted based on their socioeconomic, environmental, and demographic profile.

Therefore, there are three related objectives of our study. First, we compare two methods of LE calculation for census tracts in California: Poisson modeling and traditional life table approaches. This analysis will help us understand which method produces the most stable and valid LE estimates, as well as to understand the advantages and drawbacks of each. Although a substantial amount of research has examined these advantages and drawbacks, few have done so on a fine geographic scale.

Second, we assess the association between poverty and LE on the census tract level. This objective largely will validate previous studies that have established the negative association between poverty and LE. In this objective, we not only examine poverty, race/ethnicity, and

other demographic characteristics, we also examine the potential for interactions of these characteristics to influence life expectancy.

Third, we obtain a set of positive outlier census tracts based on residuals from the LEpoverty models described in the previous objective. These outliers represent census tracts that have unusually high life expectancy given their socioeconomic and demographic characteristics that would otherwise predict lower LE. Using those outliers, we then evaluate the associations between a detailed and comprehensive list of socioeconomic, demographic, and environmental factors and "outlier status".

Methods

OBJECTIVE 1: COMPARISON OF LIFE EXPECTANCY CALCULATION METHODS:

We compared two approaches to estimate LE at the census tract level: the actuarial life table method and Poisson modeling. We abstracted all 1999-2001 and 2009-2011 death records from California vital statistics and geocoded each record to the census tract of the decedent's last residence. Records were aggregated by age group and census tract and combined with 2000 and 2010 US Census Bureau population data, respectively. Age-specific mortality rates were calculated using average annual deaths over the three years divided by the census population count for that year. We used those rates to calculate abridged life tables and resultant censustract life expectancies at birth with both methods: the abridged life tables and Poisson modeling. Chang's method was used to calculate the variance of LE [9]. The Poisson modeling approach has three advantages for estimating LE of small area: pooling multiple year data, borrowing strengths across geospatial units, and incorporating relationships between covariates. We estimated LE by fitting mixed-effect Poisson regressions based on the same spatial and age distribution of deaths in California [10, 11]. We also compared the LEs produced by these two methods across several characteristics: consistency across point estimates, standard deviation and confidence bounds, and number of census tract estimates marked as unreliable based on extreme population sizes or tabulated deaths.

Calculation of Life Expectancy

Life Table Methods

Death records for each census tract were aggregated by 13 age groups and combined with US Census Bureau population data. Age-specific mortality rates were calculated using average annual deaths over the three years divided by the respective census population count. We used those rates to calculate abridged life tables and resultant census-tract LEs at birth detailed by the following formula (Chiang, 1984):

$$l_{x+n} = {}_{n}l_{x} - {}_{n}d_{x}, L_{0} = l_{1} + a_{0} \times d_{0} = l_{1} + a_{0} \times (l_{0} - l_{1}),$$

$${}_{n}L_{x} = \frac{n \times (l_{x} + l_{x+n})}{2}, L_{85} = \frac{l_{85}}{m_{95}}, T_{x} = \sum_{n} L_{x}, e_{x} = \frac{T_{x}}{l_{x}}, L_{1} = \frac{1}{2}$$

where x indicates the starting point for an age interval; n is the interval length and m is mortality rate; q is mortality probability; l is expected survival population; d is expected number of deaths; L is survival person-years; T is total survival person-years; E is expected remaining life time; $0 < a_0 < 1$ is a coefficient. We use infant mortality rate for 0-1 age group and the mortality probability for the last age group is set as 1. While some life tables are partitioned into one-year age groups, this study uses 13 age-groups due to the structure of the U.S. Census Bureau's American Community Survey (as accessed through the American FactFinder) allowing for reliable estimation over the broad range of subdomains considered. Assume that grouped data in the age intervals are independent across intervals and total deaths, D_i , within interval *i* follows a binomial distribution with probability q_i but with unknown number of independent trials. Chiang (1984) developed Taylor-Linearization methods to estimate the variance of LE as

$$\operatorname{var}(\hat{e}_{a}) = \sum_{i=a}^{w-1} \hat{p}_{ai}^{2} [(1-a_{i})n_{i} + \hat{e}_{i+1}]^{2} \hat{V}(\hat{p}_{i})$$

where w is the final age interval; P_{ai} is the probability of surviving to age x_i given survival to age x_a ; $q_i = 1 - p_i$, $\hat{V}(\hat{p}_i) = \hat{q}_i^2 (1 - \hat{q}_i) / D_i$.

Poisson modeling

A Poisson modeling approach was also used to estimate LE at birth, which, as previously stated, is advantageous in estimating the LE of small areas. We estimated LE by fitting mixed-effect Poisson regressions based on the same spatial and age distribution of deaths in California [2, 3]. For small areas, Poisson modeling fits better than logistic regression because the latter is biased when the outcome is rare. The model is specified as below:

$$\log(y_{ik}) = \beta_0 + \beta_1 x_{1ik} + \sum_{j=2}^{24} \beta_j x_{jik} + \sum_{j=25}^{33} \beta_j x_{jik} + z_{ik} a_k, i = 1, \dots, 6850, k = 1, \dots, 57$$

where y_{ik} is the death count within census tract i and county k.

 x_{1ik} is the poverty measure;

 x_{jik} (*j* = 2,...,11) is the population relative frequency of the *j*th age group;

 x_{jik} (*j* = 12,...,15) is the population relative frequency of the *j*th marital status group (divorced, never married, married, and separated);

 x_{16ik} is the population relative frequency of females;

 x_{jik} (j=17,18,19) is the population relative frequency of the j^{th} educational attainment group (less than high school, high school, and some college or associate degree);

 x_{20ik} is the population relative frequency of the native population;

 x_{21ik} is the population relative frequency of Hispanic ethnicity.

The parameter β_0 is an intercept term, $\beta_i (i = 1, \dots, 21)$ is the respective slope of the $x_i (i = 1, \dots, 21)$; Z_{ik} is the random effect for the ith census tract and jth county in California; and a_k is the random coefficient effect for the kth county.

OBJECTIVE 2: SMALL-AREA LIFE EXPECTANCY AND POVERTY

Life table methods were ultimately selected and were used to calculate the LEs of census tracts with a population of at least 3750 individuals, \geq 20 deaths, and at least 50% of population not living group quarters such as correctional institutions, nursing home, college dormitories including college quarters off campus, or military quarters. This results in a sample of 6,670 and 7,654 census tracts used in this analysis for 2000 and 2010, respectively. A parsimonious weighted least squares (WLS) regression (model 1) was selected among the following covariates: poverty, education, age, gender, marital status, race, immigrant status, and Hispanic origin from the US Census using reciprocal variance of the life expectancy as regression weights. The variables were selected *a priori* as major predictors of life expectancy in previous studies.

Small-area life expectancy, poverty, and other demographic characteristics

Weighted least squares regression

The dependent variable (LE) is continuous and census tracts have differential variability in LE. Therefore, a WLS is suitable for analysis of such data. The maximum likelihood estimates of the linear component parameters are calculated using SAS/STAT®, version 9.4, PROC GLM. ArcMap v. 10.1 (ESRI, Redlands, CA) was used for mapping.

The model for a given census tract is

$$y_{i} = \beta_{0} + \beta_{1}x_{1} + \sum_{j=2}^{24}\beta_{j}x_{j} + \sum_{j=25}^{33}\beta_{j}x_{j} + \varepsilon_{i}, \operatorname{var}(\varepsilon_{i}) \Box N(0, \frac{\sigma^{2}}{w_{i}}), i = 1, \cdots, 6850,$$

where w_i (i=1,...,6850) is the standard deviation of LE for the ith census tract in California;

- x_1 is the poverty measure;
- x_j (*j* = 2,...,11) is the population relative frequency of the *j*th age group;
- x_j (*j* = 12,...,15) is the population relative frequency of the *j*th marital status group (divorced, never married, married, and separated);
- x_{16} is the population relative frequency of females;
- x_j (j=17,18,19) is the population relative frequency of the j^{th} educational attainment group (less than high school, high school, and some college or associate degree);
- x_{20} is the population relative frequency of the native population;

 x_{21} is the population relative frequency of Hispanic ethnicity.

The parameter β_0 is an intercept term, $\beta_i (i = 1, \dots, 21)$ is the respective slopes of the $x_i (i = 1, \dots, 21)$.

Associations between life expectancy and poverty

Poverty was measured as the proportion of persons with a total household income below 200% of the federal poverty level. LE was negatively associated with poverty with a correlation of -0.47. The association is illustrated in Figure 1. The scatterplot was smoothed using a 30-point moving average after truncating 15 points on each side, left and right. The relationship between LE and poverty was not linear as would be expected, with the mean LE (approximated by the moving average) plateauing at a high LE on the left and leveling out at a low LE on the right.

OBJECTIVE 3: OUTLIER ANALYSIS

The analysis then examined outlier census tracts to ascertain those community-level assets that may contribute to unexpectedly high life expectancy even after controlling for poverty. In this phase of the analysis, census-tract life expectancies were categorized as three classes: unexpectedly low LE with a standardized residual value less than or equal to -2.576, unexpectedly high LE with a standardized residual value greater than or equal to 2.576, and expectedly normal LE with an absolute standardized residual value between -2.576 and 2.576. We then used an extensive list of predictor variables from the US Census, and numerous other state and national databases on socioeconomic, environmental, demographic, and health-related factors (Appendix A). A weighted multinomial logistic regression was modeled against each predictor controlling for the significant variables in model 1. WLS is an effective method to make good use of small-area data sets and can handle data points with varying quality in regression analysis. Weights used in the analysis were the reciprocal of the variance associated with the LE calculated in the Phase 1 regression models.

Results

In this section, we first evaluate models 1 and 2 and then summarize the results from these models. We also compare the results from model 1 with those from model 2. Finally, we document the results of outlier analysis.

Objective 1: Comparison of life expectancy calculation methods

The life expectancies produced by these two methods were compared across several characteristics: consistency across point estimates, standard deviation and confidence bounds, number of census tract estimates marked as unreliable based on extreme population sizes or tabulated deaths, and adjusted R-squared values. Summary statistics for both methods and their comparison are shown in Tables 1 and 2. The results, and the respective abilities of the models to predict life expectancy, suggest that the life table approach is more reliable than Poisson modeling. Census tract-specific life expectancies for 2010 are displayed in Figure 1 using this approach.

Figures 2A and 2B indicate that the scatter plots for observed LE against predicted LE and the difference between observed LE and predicted LE by 13 age groups using Poisson modeling. Figure 3A indicates that predicted LE from Poisson modeling overestimates the observed LE. The distribution of the differences between observed LE and predicted LE are centered at zero for the first eleven age groups, but spreads out for the last two age groups. It appears there are some potential outliers in the last two age groups.

Based on these results, the subsequent analyses will be conducted using the life table approach for calculating LE.

Objective 2: Small-area life expectancy and poverty ("Phase 1")

The results showed a monotonic negative association: life expectancy decreased as the population with incomes below 200% of the Federal poverty threshold increased, and this association held even after adjusting for confounders. Overall, 71% of the variability in life expectancy was explained by the Phase 1 variables (Table 3). Scatterplots are shown to illustrate how outliers were identified (census tracts with large standardized residuals that depart from the regression equation) and eliminated outliers suspected to be erroneous based on implausible LE values (Figure 4).

Objective 3: Outlier analysis ("Phase 2")

We identified a listing of area-based measures (Appendix A) with the potential to account for some of the outlier census tract's higher than expected life expectancies, including environmental, socioeconomic, detailed demographic, and health-related variables, beyond the variables used in the Phase 1 model. Some variables are modifiable, while others are nonmodifiable, including many demographic variables. Recall that weighted multinomial logistic regression was used to model outlier status as a function of these variables.

Prediction of Negative Outliers (Low LE)

Table 4 displays the frequencies of both positive and negative outliers for data 2000 and 2010. Some of the strongest predictors of negative outliers (unexpectedly low LE given low poverty levels) were the ratio of whites to non-whites (odds ratio (OR) 0.85, 90% CI [0.72, 0.98]), and percent of population living in nursing homes (OR 23.4, 90% CI 7.84, 71.4), although census tracts with a majority of nursing home residents were removed a priori from the analysis. Positive predictors were less common, however.

Prediction of Positive Outliers (High LE)

Much of the remainder of the analysis focuses on community assets that predicted positive outliers- those census tracts with unexpectedly high life expectancy after controlling for the basic demographic and socioeconomic variables in Phase I. First, using weighted logistic regression, we regressed positive outlier status on each of the "Phase 2" variables individually in what we define as the "bivariate analysis". Second, we conducted a multivariable logistic regression analysis of those outliers including all significant predictors (p < 0.2) from the bivariate analysis. The bivariate and multivariate analyses were conducted on the 2000 and 2010 data separately.

Tables 5 and 6 represent bivariate of outliers for 2000 and 2010, respectively. For the 2000 analysis, some of the strongest associations between predictors and positive outliers (unexpectedly high LE) included percent of population who bikes or walks to work (OR 1.85, 90% CI: 1.19, 2.69). This result means that, controlling the phase 1 variables, a 10 percent increase in the population who bike or walk to work will increase the likelihood of being an outlier census tract (high LE, high poverty) by 85%. Factors that significantly reduced the likelihood of outlier status included cancer risk due to air pollutants (OR 0.76, 90% CI: 0.68, 0.84), foster care rate (OR 0.06, 90% CI: 0.02, 0.25), percent of householders without children

(OR 0.59, 90% CI: 0.35, 0.99), and percent of households that were overcrowded (>= 1.01 persons/room, OR 0.45, 90% CI: 0.19, 1.00). Factors that were positively associated with outlier status included percent population living in rural areas (OR 1.23, 90% CI: 1.14, 1.33) and sex ratio (female to male) at birth (OR 1.63, 90% CI: 1.04, 2.49).

Similarly, for 2010 data some of the strongest associations between predictors and positive outliers (unexpectedly high LE) included single-parent families (OR 0.20, 90% CI: 0.11, 0.33), dual parent and single father families (OR 2.42, 90% CI: 1.69, 3.48), annual mean PM 2.5 concentrations (OR 0.54, 90% CI: 0.39, 0.75), illiteracy rate (OR 0.71, 90%CI: 0.54, 0.93), presence of college dormitories including college quarters off campus (OR 1.04, 90% CI: 1.01, 1.07), the presence of an acute care hospital (OR 2.60, 90% CI: 1.48, 4.31), The presence of a specialty clinic (OR 2.31, 90% CI: 1.32, 3.78), CT is predominantly open space (OR 2.52, 90% CI: 1.45, 4.12), and percent of the population living within 1/2 mile of a park, beach, open space, or coastline (OR 0.95, 90% CI: 0.90, 1.00).

The results of multiple logistic regression models are summarized in Tables 7 and 8 for data 2000 and 2010, respectively. After adjustment, several factors remained significant in 2000. Census tracts considered rural were about 19% more likely to be a positive outlier than more urban census tracts (OR 1.19, 90% CI: 1.08, 1.31), for example. Negative predictors of outlier status included cancer risk due to air pollutants (OR 0.65, 90% CI: 0.51, 0.84) and low travel time to work: 30 minutes or less (OR 0.65, 90% CI: 0.48, 0.86) and 30-60 minutes (OR 0.52, 90% CI: 0.35, 0.77). In 2010, the model was more robust, yielding many more significant factors associated with outlier status. In addition to several family structure variables, being a high LE/high poverty outlier was associated with the presence of at least one acute care hospital (OR 2.11, 90% CI: 1.08, 4.12) and at least one specialty clinic (OR 2.06, 90% CI: 1.07, 3.96) in

that census tract. Higher inequality, as measured through the Gini coefficient, (OR 1.67, 90% 1.12, 2.49) was also associated with an increased likelihood of being an outlier. Two environmental variables were also predictive of outlier status: open space was positively associated (OR 2.55, 90% CI: 1.39, 4.66) and PM 2.5 concentrations were negatively associated (OR 0.88, 90% CI: 0.80, 0.97).

Discussion

There were numerous options available in selecting the most appropriate manner in which to calculate LE, and several strengths and drawbacks of each method. Despite these limitations, the results of Objective 1 show that reliable LE estimation is possible when properly applied to small geographic areas, with several important caveats. One such caveat is that in the analysis each census tract was considered to be independent of the others. Therefore, spatial autocorrelation that may account for some of the observed patterns was not taken into account. Second, as with any small area estimate of population health, the predicted point estimates of LE were subject to statistical uncertainty. In addition, deaths were counted at the address of the decedent's last known residence. Therefore, the factors that may contribute to life expectancy may come from other census tracts in which decedents may work or otherwise spend more time. Lastly, there were numerous counterintuitive findings, particularly in the multivariate models (Tables 7 and 8) for both years. In the multivariate models, it is possible that highly correlated predictors may result in one predictor being a "positive" predictor of outlier status while the other may emerge as an unexpected "negative" predictor to compensate statistically. An example of this potential limitation may have occurred in 2000, where the maltreatment allegation rate was actually a positive predictor of outlier status. This unexpected finding may be due to collinearity with other, similar variables in the model that

may be associated with outlier status in the direction we would expect. Other examples include unexpected findings regarding income inequality, commuting time to work, and proximity to parks, beaches, and other open spaces.

In Objective 2, our findings validate previous studies that have shown consistent associations between summary measures of population health—namely LE—and socioeconomic status, as measured by poverty. In Objective 3, we found additional, less well- established associations that might account for some of the extreme values of LE observed above and beyond the contribution of poverty to explain LE. Our preliminary results suggest that factors that may help communities enjoy longer life expectancy include cleaner air, access to health care facilities, employment, and literacy. Such findings may be useful to policymakers who cannot rectify core social determinants such as household income or education, but are able to institute policies that may buffer their adverse effects. For example, the finding that census tracts with high illiteracy rates were significantly less likely to be a positive outlier could have potential importance for policymakers. Some factors that were also found to predict outlier status, such as rural locations, open space, and the prevalence of single-parent households are less amenable to change by policymakers, however. The next step in the analysis is to delve deeper into assessing those census tracts that remain outliers after accounting for other socioeconomic, environmental, and demographic factors in Objective 3. These remaining census tracts will now undergo quantitative spatial analysis and qualitative inquiry (key informant interviews) to further clarify positive assets that might explain their more favorable outcomes.

These positive assets may influence health directly or indirectly through facilitating or impeding healthy behaviors. Although unhealthy behavior is a matter of personal choice to some degree, the influence of neighborhood and community factors on health behavior and population

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health is well-documented [12-14]. Personal health behaviors, such as eating a healthy diet and actively exercising, are intrinsically tied to socioeconomic and social and physical environment factors, such as some of those assessed in this analysis. Still other community assets are not determinative of behaviors.

The debate over individual versus area-level factors and their influence on population health remains unresolved. The findings of this study underscore the potential for communitybased assets to affect health, above and beyond the contributions of traditional socioeconomic factors, such as income and poverty. Moreover, some community-level factors are more than simply aggregated individual-level factors and could not be measured on the individual level, particularly environmental and related factors.

Our findings highlight the utility, strengths, and drawbacks of life expectancy methods in these small geographic areas and validate the myriad of past studies on the association between population health and socioeconomic status. In this study we also present a novel way of classifying outliers from the poverty-life expectancy association and assessing the neighborhood and environmental characteristics that may contribute to unexpectedly high or low life expectancy. Assessing variation and ultimately reducing geographic disparities in population health is a major challenge today. Our study examines several potentially modifiable factors that can improve health, save lives, and reduce disparities through evidence-supported policy.

What this study adds:

- A comparison of life expectancy calculation methods for small geographic areas
- A description of some of the strengths and drawbacks to small-area life expectancy calculation

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- A novel statistical method for categorizing and describing small areas with unexpectedly high life expectancy and examining predictors of those outliers
- A detailed assessment of small area characteristics that potentially predict life expectancy

Table 1: Summary statistics of life expectancy (LE) from two approaches, 2010

| Type of LE calculation | Ν | mean | SE | min. | median | max. |
|------------------------|------|-------|------|-------|--------|--------|
| Observed LE | 6435 | 77.93 | 3.67 | 60.82 | 78.03 | 98.51 |
| LE from Poisson | 6435 | 80.69 | 3.29 | 66.30 | 80.62 | 122.81 |

| Population (low bound) | Ν | Mean | SE | Median |
|-------------------------------|------|------|------|--------|
| 951 | 94 | 2.1 | 6.36 | 2.03 |
| 5000 | 374 | 2.52 | 5.37 | 2.39 |
| 7500 | 915 | 2.63 | 4.63 | 2.58 |
| 10000 | 2477 | 2.67 | 4.05 | 2.43 |
| 15000 | 1762 | 2.76 | 3.86 | 2.37 |
| 20,000 | 801 | 2.98 | 3.57 | 2.69 |
| 25,000 | 320 | 3.91 | 4.04 | 3.61 |

 Table 2: Comparison of life expectancy (LE) using three approaches, 2010

| Variable | Parameter estimate (beta) | Standard error | T-score | P-value |
|----------------------------|---------------------------|----------------|----------------|---------|
| Intercept | 75.45816 | 1.73755 | 43.43 | <.0001 |
| Poverty | | | | |
| Below 200% of FPL | 0.17736 | 0.01883 | 9.42 | <.0001 |
| Poverty squared | -0.000762 | 0.000163 | -4.68 | <.0001 |
| Sex | | | | |
| Percent female | 0.02045 | 0.01067 | 1.92 | 0.0553 |
| Age | | | | |
| Percent < 5 years old | -0.03055 | 0.01651 | -1.85 | 0.0643 |
| Percent 5 to 9 years old | 0.06367 | 0.03108 | 2.05 | 0.0406 |
| Percent 20 to 24 years old | 0.08159 | 0.01427 | 5.72 | <.0001 |
| Percent 60 to 64 years old | 0.06867 | 0.01934 | 3.55 | 0.0004 |
| Percent 65 to 74 years old | 0.07746 | 0.01661 | 4.66 | <.0001 |
| Percent 75 to 84 years old | 0.13465 | 0.01924 | 7.00 | <.0001 |
| Marital status | | | | |
| Percent divorced | 0.13874 | 0.02434 | 5.70 | <.000 |
| Percent never married | 0.12149 | 0.01811 | 6.71 | <.000 |
| Percent married | 0.12832 | 0.01434 | 8.95 | <.0001 |
| Education | | | | |
| Percent with no high | -0.11769 | 0.00757 | -15.54 | <.0001 |
| school degree | | | | |
| Percent with high school | -0.07456 | 0.01121 | -6.65 | <.000 |
| diploma only | | | | |
| Percent with some college | -0.08622 | 0.00600 | -14.37 | <.000 |
| and associates degree | | | | |
| Ethnicity | | | | |
| Percent Native-born | -0.02546 | 0.00567 | -4.49 | <.000 |
| Percent Hispanic | 0.00901 | 0.00572 | 1.58 | 0.1150 |
| Interactions | | | | |
| poverty*age2 | -0.00215 | 0.000929 | -2.32 | 0.0200 |
| poverty*educ2 | -0.00141 | 0.000358 | -3.93 | <.000 |
| poverty*Divorced | -0.00289 | 0.000629 | -4.60 | <.000 |
| poverty*Never_married | -0.00131 | 0.000293 | -4.47 | <.000 |
| poverty*native | -0.00150 | 0.000176 | -8.46 | <.000 |
| poverty*Hispanic | 0.00050 | 0.000139 | 3.57 | 0.0004 |

Table 3: Results of Phase 1 analysis "Best model": Predictors of life expectancy based on
poverty and other community-level sociodemographic characteristics for 2010

Table 4: Frequency of positive and negative outliers, 2000 and 2010. Positive outliers are those census tracts that had unexpectedly high LE given high poverty, whereas negative outliers are those census tracts with unexpectedly low life expectancy and low poverty

| | 2000 | | 2010 | |
|--------------|-----------|---------|-----------|---------|
| Outlier type | Frequency | Percent | Frequency | Percent |
| -1 | 50 | 0.75 | 86 | 1.12 |
| 0 | 6538 | 98.02 | 7491 | 97.87 |
| 1 | 82 | 1.23 | 77 | 1.01 |

| Variable | Label | Unit | | Life ex | spectancy | (1) |
|-------------|---|------|---------------|---------|-----------|---------|
| | | | Odds Ratio | 90% | 6 CI | p-value |
| Ainsured | percent of adults insured | 10.0 | 1.00 | 0.62 | 1.63 | 0.9868 |
| Allinsured | percent of all age insured | 10.0 | 1.01 | 0.60 | 1.71 | 0.9762 |
| Allrate | maltreatment allegation rate(per 1000 children) | 10.0 | 0.96 | 0.85 | 1.09 | 0.6268 |
| Armed_force | Work in the military | 10.0 | 0.41 | 0.05 | 1.71 | 0.4338 |
| BikWaL | Walk or bike to work | 10.0 | 1.85 | 1.19 | 2.69 | 0.0114 |
| Cancer | Cancer risk due to air pollutants (per million) | 10.0 | 0.76 | 0.68 | 0.84 | <.0001 |
| Employed | work but not in the military | 10.0 | 0.65 | 0.42 | 1.08 | 0.1335 |
| Entrrate | enter foster care rate(per 1000 children) | 10.0 | 0.06 | 0.02 | 0.25 | 0.0010 |
| FFR | Number of Fast Food Restaurants | 10.0 | 1.22 | 0.69 | 1.96 | 0.5414 |
| Fami11 | Two-parent family householders | 10.0 | 1.23 | 0.69 | 2.16 | 0.5531 |
| Fami12 | single-parent family householders | 10.0 | 0.63 | 0.25 | 1.42 | 0.3884 |
| Fami13 | no children family householders | 10.0 | 0.59 | 0.35 | 0.99 | 0.0874 |
| Fami14 | non-family householders | 10.0 | 1.41 | 0.93 | 2.15 | 0.1838 |
| Fami21 | Two-parent families | 10.0 | 1.71 | 0.94 | 3.04 | 0.1340 |
| Fami22 | single-parent families | 10.0 | 0.87 | 0.44 | 1.66 | 0.7335 |
| Fami23 | no children families | 10.0 | 0.63 | 0.35 | 1.17 | 0.2104 |
| Fami31 | Single mother families | 10.0 | 0.96 | 0.62 | 1.49 | 0.8950 |
| Fami32 | Single father or two-parent families | 10.0 | 1.04 | 0.67 | 1.63 | 0.8935 |
| GINI | | 0.1 | 2.82 | 1.47 | 5.41 | 0.0088 |
| HFP | % health food | 10.0 | 1.00 | 0.83 | 1.23 | 0.9999 |
| Immun | Percent of kindergarteners with all immunizations | 0.5 | 0.93 | 0.03 | 32.8 | 0.9712 |
| Kinsured | percent of under age 18 insured | 10.0 | 1.01 | 0.56 | 1.84 | 0.9800 |
| MHPR | Mental public health provider ratio | 10.0 | 0.92 | 0.84 | 1.01 | 0.1565 |

Table 5: Results of marginal weighted logistic regression analysis adjusted for Phase 1variables (2000)

| Neuro | Neurological risk due to air | 1.0 | 0.00 | 0.00 | 0.02 | <.0001 |
|-----------|--|------|------|------|------|----------|
| DCDD | pollutants (HQ) | 10.0 | 0.02 | 0.04 | 1.01 | 0 1 40 2 |
| PCPR | PCP rate | 10.0 | 0.92 | 0.84 | 1.01 | 0.1483 |
| PM25_00 | PM2.5 in 2000 | 10.0 | 1.02 | 0.86 | 1.19 | 0.8767 |
| PUnheal | Percent of at least unhealthy days | 10.0 | 0.99 | 0.88 | 1.11 | 0.8570 |
| Quart1 | correctional institutions | 10.0 | 0.58 | 0.07 | 1.44 | 0.4872 |
| Quart2 | nursing homes | 10.0 | 0.16 | 0.01 | 1.94 | 0.3065 |
| Quart3 | college dormitories including college quarters off campus | 10.0 | 2.07 | 0.65 | 4.89 | 0.2001 |
| Quart5 | other institutional and no institutional group quarters | 10.0 | 2.21 | 1.10 | 4.24 | 0.0510 |
| Quart6 | | 10.0 | 0.72 | 0.45 | 1.21 | 0.2645 |
| QuartP | either group quarters present or absent | 1.0 | 0.96 | 0.63 | 1.49 | 0.8701 |
| RFR | % recreation facility rate | 10.0 | 1.40 | 0.79 | 2.39 | 0.3185 |
| RSega | the ratio between White and nonwhite | 1.0 | 1.02 | 0.95 | 1.09 | 0.5731 |
| Race1 | White | 10.0 | 1.01 | 0.85 | 1.21 | 0.9295 |
| Race2 | Black | 10.0 | 0.69 | 0.44 | 0.97 | 0.1158 |
| Race3 | American Indian, Eskimo, or Aleut | 10.0 | 1.82 | 0.53 | 4.08 | 0.2375 |
| Race4 | Asian and Pacific Islander | 10.0 | 1.19 | 0.92 | 1.56 | 0.2704 |
| Race5 | Other | 10.0 | 1.22 | 0.64 | 2.34 | 0.6195 |
| RentMort | burden and overburden for Rent or mortgage | 10.0 | 0.97 | 0.69 | 1.36 | 0.8696 |
| RentMort1 | burden and overburden for Rent or mortgage excluding the no computed units | 10.0 | 0.89 | 0.63 | 1.26 | 0.5776 |
| Respir | Respiratory risk due to air pollutants (HQ) | 1.0 | 0.93 | 0.90 | 0.97 | 0.0011 |
| RoomO1 | % of households in overcrowded (>= 1.01 persons/room) | 10.0 | 0.45 | 0.19 | 1.00 | 0.1084 |
| RoomO2 | severely overcrowded (>= 1.5 persons per room) | 10.0 | 1.11 | 0.65 | 1.87 | 0.7529 |
| Rural | | 10.0 | 1.23 | 1.14 | 1.33 | <.0001 |
| Sexratio | Sex ratio at birth | 0.1 | 1.63 | 1.04 | 2.49 | 0.0676 |

| Subsrate | substantiation rate(per 1000 children) | 5.0 | 0.67 | 0.53 | 0.84 | 0.0039 |
|------------|---|------|------|------|------|--------|
| TravT1 | percent of population(not work at home)to travel less than 30 minutes | 10.0 | 1.02 | 0.87 | 1.20 | 0.8277 |
| TravT2 | percent of population(not work at home)to travel 30 to 60 minutes | 10.0 | 0.75 | 0.60 | 0.93 | 0.0300 |
| TravT3 | percent of population(not work at home)to travel more than 60 minutes | 10.0 | 1.50 | 1.15 | 1.94 | 0.0102 |
| Unemployed | | 10.0 | 1.95 | 1.12 | 3.28 | 0.0399 |
| Urban | | 10.0 | 0.81 | 0.75 | 0.88 | <.0001 |
| illiterate | percent of illiterate population | 10.0 | 0.71 | 0.54 | 0.93 | 0.0395 |

| | | | |] | LE(1) | |
|-------------|---|------|---------------|------|-------|--------|
| Variable | Label | Unit | Odds ratio | L | U | Prob |
| Ainsured | percent of adults (19-64) insured | 10.0 | 0.97 | 0.72 | 1.30 | 0.8470 |
| Allinsured | percent of all age(under 65) insured | 10.0 | 0.94 | 0.64 | 1.38 | 0.7866 |
| Allrate | maltreatment allegation rate (per 1000 children) | 1.0 | 0.99 | 0.98 | 1.00 | 0.1418 |
| Armed_force | | 10.0 | 1.26 | 0.40 | 2.56 | 0.6852 |
| BikWaL | Walk or bike to work | 10.0 | 1.38 | 0.94 | 1.94 | 0.1437 |
| Cinsured | Percent of children insured | 10.0 | 0.00 | 0.00 | 1E25 | 0.6616 |
| Diesel_PM | Diesel PM emissions from on-road and non-road sources | 10.0 | 0.80 | 0.67 | 0.95 | 0.0358 |
| Employed | Percent employed | 10.0 | 1.20 | 0.77 | 1.92 | 0.5038 |
| Entrrate | enter foster care rate(per 1000 children) | 10.0 | 0.43 | 0.12 | 1.54 | 0.2880 |
| Fami11 | Two-parent family householders | 5.0 | 1.25 | 1.06 | 1.47 | 0.0280 |
| Fami12 | single-parent family householders | 10.0 | 0.16 | 0.08 | 0.32 | <.0001 |
| Fami13 | no children family householders | 10.0 | 1.26 | 0.86 | 1.83 | 0.3246 |
| Fami14 | non-family householders | 10.0 | 0.91 | 0.71 | 1.15 | 0.4976 |
| Fami21 | Two-parent families | 10.0 | 2.08 | 1.44 | 3.00 | 0.0010 |
| Fami22 | single-parent families | 10.0 | 0.20 | 0.11 | 0.33 | <.0001 |
| Fami23 | no children families | 10.0 | 1.12 | 0.78 | 1.58 | 0.6110 |
| Fami31 | Single mother families | 10.0 | 0.41 | 0.29 | 0.59 | <.0001 |
| Fami32 | Single father or two-parent families | 10.0 | 2.42 | 1.69 | 3.48 | <.0001 |
| Immun | Percent of kindergarteners with all immunizations | 0.1 | 1.00 | 0.99 | 1.00 | 0.4793 |

Table 6: Results of marginal weighted logistic regression adjusted for Phase 1 variables(2010)

| Kinsured | percent of under age 18 insured | 10.0 | 0.46 | 0.18 | 1.13 | 0.1584 |
|------------|--|------|------|------|------|--------|
| PM25 | annual mean PM 2.5 concentrations | 5.0 | 0.54 | 0.39 | 0.75 | 0.0019 |
| PM25_10 | PM2.5 in 2010 | 2.0 | 0.93 | 0.84 | 1.02 | 0.2496 |
| PUnheal | Percent of at least unhealthy days | 1.0 | 0.99 | 0.98 | 1.01 | 0.4103 |
| Pesticides | Total pounds of selected active pesticide ingredients used in production-agriculture | 10.0 | 1.00 | 1.00 | 1.00 | 0.0137 |
| Quart2 | nursing homes | 0.1 | 0.99 | 0.97 | 1.01 | 0.3862 |
| Quart3 | college dormitories including college quarters off campus | 0.5 | 1.04 | 1.01 | 1.07 | 0.0059 |
| Quart5 | other institutional and non-institutional group quarters | 10.0 | 0.61 | 0.11 | 2.03 | 0.5880 |
| QuartP | either group quarters present or absent | 1.0 | 1.22 | 0.82 | 1.86 | 0.4235 |
| RSega | the ratio between White and nonwhite | 1.0 | 0.99 | 0.93 | 1.03 | 0.7509 |
| RentMort | burden and overburden for Rent or mortgage | 10.0 | 1.17 | 0.95 | 1.43 | 0.2075 |
| RoomO1 | % of households in overcrowded (>= 1.01 persons/room) | 10.0 | 1.27 | 0.75 | 2.10 | 0.4493 |
| RoomO2 | severely overcrowded (>= 1.5 persons per room) | 0.5 | 1.00 | 0.97 | 1.03 | 0.8840 |
| Rural | | 10.0 | 0.98 | 0.86 | 1.09 | 0.7616 |
| Sexratio | Sex ratio at birth | 1.0 | 0.28 | 0.00 | 117 | 0.7505 |
| Subsrate | substantiation rate(per 1000 children) | 10.0 | 0.96 | 0.62 | 1.46 | 0.8749 |
| TravT1 | percent of population(not work at home)to travel less than 30 minutes | 10.0 | 1.06 | 0.93 | 1.21 | 0.4402 |
| TravT2 | percent of population(not work at home)to travel 30 to 60 minutes | 10.0 | 0.86 | 0.73 | 1.03 | 0.1630 |
| TravT3 | percent of population(not work at home)to travel more than 60 minutes | 1.0 | 1.01 | 0.98 | 1.03 | 0.5751 |
| Unemployed | Percent unemployed | 10.0 | 0.77 | 0.47 | 1.24 | 0.3704 |
| | | | | | | |

| Urban | | 10.0 | 1.02 | 0.92 | 1.16 | 0.7616 |
|---------|---|------|------|------|------|--------|
| bed3 | Presence of beds for inpatient mental health (versus absent) | 1.0 | 3.03 | 0.71 | 8.56 | 0.1280 |
| bed5 | Presence of beds for long-term care facilities either (versus absent) | 1.0 | 0.90 | 0.51 | 1.48 | 0.7308 |
| bed7 | Presence of beds for acute care hospitals (versus absent) | 1.0 | 2.60 | 1.48 | 4.31 | 0.0031 |
| cancer | Cancer risk due to air pollutants (per million) | 0.1 | 0.00 | 0.00 | 2E30 | 0.1483 |
| cost1 | child care center infant cost | 10.0 | 1.00 | 1.00 | 1.00 | 0.3001 |
| cost2 | child care center preschooler cost | 10.0 | 1.00 | 1.00 | 1.00 | 0.2877 |
| cost3 | family child care home infant cost | 0.1 | 1.00 | 1.00 | 1.00 | 0.2292 |
| cost4 | family child care home preschooler cost | 10.0 | 1.00 | 1.00 | 1.00 | 0.3154 |
| develop | area is developed or not | 1.0 | 1.52 | 0.96 | 2.51 | 0.1500 |
| enrol1 | Percent of age group enrolled in school | 10.0 | 1.00 | 0.94 | 1.07 | 0.9906 |
| enrol2 | Percent of age group enrolled in public school | 1.0 | 1.00 | 0.99 | 1.01 | 0.8666 |
| enrol3 | Percent of age group enrolled in private school | 10.0 | 1.01 | 0.95 | 1.06 | 0.8659 |
| facip1 | specialty clinics either present or absent | 1.0 | 2.31 | 1.32 | 3.78 | 0.0084 |
| facip3 | inpatient mental health either present or absent | 1.0 | 3.03 | 0.71 | 8.56 | 0.1280 |
| facip4 | community/free clinic either present or absent | 1.0 | 1.08 | 0.60 | 1.83 | 0.8103 |
| facip5 | long-term care facilities either present or absent | 1.0 | 0.90 | 0.51 | 1.48 | 0.7308 |
| facip6 | home health/hospice either present or absent | 1.0 | 0.81 | 0.46 | 1.33 | 0.5108 |
| facip7 | acute care hospitals either present or absent | 1.0 | 2.60 | 1.48 | 4.31 | 0.0031 |
| farm | area is cultivated or not | 1.0 | 0.36 | 0.10 | 0.92 | 0.1150 |
| | | | | | | |

| gini | Gini coefficient | 10.0 | 4E26 | 2E12 | 2E40 | 0.0018 |
|----------------|---|------|------|------|------|--------|
| lead1 | 0-5 years with elevated blood lead level | 1.0 | 26E6 | 0.00 | 9E48 | 0.7820 |
| lead2 | 6-20 years with elevated blood lead level | 1.0 | 0.00 | 0.00 | 7E6 | 0.4029 |
| lead3 | all children under 20 with elevated blood level | 1.0 | 17E3 | 0.00 | 1E45 | 0.8741 |
| neuro | Neurological risk due to air pollutants (HI) | 0.1 | 0.86 | 0.50 | 1.38 | 0.6253 |
| openspace | area is open space versus not | 1.0 | 2.52 | 1.45 | 4.12 | 0.0034 |
| park | Percent of population within 1/2 mile of park, beach, open space, or coastline | 10.0 | 0.95 | 0.90 | 1.00 | 0.0825 |
| pollution_burd | Average of percentiles from the pollution burden indicators | 10.0 | 0.95 | 0.83 | 1.09 | 0.5138 |
| race1 | White | 10.0 | 1.00 | 0.89 | 1.13 | 0.700 |
| race2 | Black | 10.0 | 0.41 | 0.25 | 0.64 | 0.0018 |
| race3 | American Indian, Eskimo, or Aleut | 10.0 | 1.15 | 0.27 | 2.84 | 0.8543 |
| race4 | Asian and Pacific Islander | 10.0 | 1.22 | 1.00 | 1.49 | 0.0998 |
| race5 | Other | 10.0 | 1.10 | 0.88 | 1.38 | 0.4816 |
| respir | Respiratory risk due to air pollutants (HI) | 10.0 | 0.86 | 0.28 | 2.47 | 0.8260 |
| snap | percent of household receiving snap | 10.0 | 0.58 | 0.34 | 0.96 | 0.0856 |
| traffic | Traffic density in vehicle kilometers per hour per road length within 150 meters | 10.0 | 1.00 | 1.00 | 1.00 | 0.5889 |

| Effect | Label | Unit | nit LE(1) | | E(1) | | | |
|----------|--|------|---------------|------|------|---------|--|--|
| | | | Odds Ratio | 90% | O CI | p-value | | |
| Allrate | maltreatment allegation rate (per 1000 children) | 10.0 | 1.39 | 1.11 | 1.74 | 0.02 | | |
| BikWaL | Walk or bike to work | 10.0 | 1.15 | 0.62 | 2.15 | 0.71 | | |
| Cancer | Cancer risk due to air pollutants (per million) | 10.0 | 0.65 | 0.51 | 0.84 | <0.001 | | |
| Employed | work but not in the military | 1.0 | 1.07 | 0.89 | 1.29 | 0.54 | | |
| Entrrate | enter foster care rate(per 1000 children) | 1.0 | 0.63 | 0.51 | 0.78 | <0.001 | | |
| Fami13 | no children family householders | 10.0 | 0.96 | 0.19 | 4.93 | 0.97 | | |
| Fami14 | non-family householders | 10.0 | 0.93 | 0.34 | 2.53 | 0.90 | | |
| Fami21 | Two-parent families | 10.0 | 0.59 | 0.24 | 1.44 | 0.33 | | |
| Fami23 | no children families | 10.0 | 0.33 | 0.08 | 1.25 | 0.1′ | | |
| GINI | | 0.1 | 1.26 | 0.62 | 2.56 | 0.6 | | |
| Immun | Percent of kindergarteners with all immunizations | 0.1 | 0.83 | 0.56 | 1.21 | 0.4 | | |
| MHPR | Mental public health provider ratio | 10.0 | 0.87 | 0.76 | 0.99 | 0.0 | | |
| Neuro | Neurological risk due to air pollutants (HQ) | 0.1 | 1.05 | 0.71 | 1.54 | 0.84 | | |
| PCPR | PCP rate | 10.0 | 1.13 | 0.96 | 1.32 | 0.2 | | |
| Quart1 | correctional institutions | 1.0 | 0.96 | 0.80 | 1.15 | 0.69 | | |
| Quart2 | nursing homes | 1.0 | 0.77 | 0.57 | 1.06 | 0.13 | | |
| Quart5 | other institutional and non- institutional group quarters | 1.0 | 1.04 | 0.91 | 1.19 | 0.64 | | |
| Quart6 | | 1.0 | 0.96 | 0.85 | 1.09 | 0.6 | | |
| QuartP | either group quarters present or absent | 1.0 | 0.91 | 0.59 | 1.41 | 0.74 | | |
| RFR | % recreation facility rate | 10.0 | 0.36 | 0.12 | 1.04 | 0.1 | | |
| RSega | the ratio between White and nonwhite | 1.0 | 0.92 | 0.83 | 1.01 | 0.1 | | |
| Race2 | Black | 10.0 | 0.64 | 0.40 | 1.03 | 0.12 | | |
| Respir | Respiratory risk due to air pollutants (HQ) | 1.0 | 1.08 | 1.02 | 1.15 | 0.04 | | |
| RoomO1 | % of households in overcrowded (>= 1.01 persons/room) | 10.0 | 0.58 | 0.24 | 1.37 | 0.30 | | |

Table 7: Phase-II Results from "best" multiple logistic regression model controlling for phase-I variables (2000)

| RoomO2 | severely overcrowded (>= 1.5 persons per room) | 10.0 | 1.09 | 0.64 | 1.87 | 0.79 |
|------------|---|------|------|------|------|------|
| Rural | | 10.0 | 1.19 | 1.08 | 1.31 | 0.00 |
| Sexratio | Sex ratio at birth | 0.0 | 1.04 | 1.00 | 1.08 | 0.09 |
| Subsrate | substantiation rate(per 1000 children) | 5.0 | 0.84 | 0.60 | 1.17 | 0.39 |
| TravT1 | percent of population(not work at home)to travel less than 30 minutes | 10.0 | 0.65 | 0.48 | 0.86 | 0.01 |
| TravT2 | percent of population(not work at home)to travel 30 to 60 minutes | 10.0 | 0.52 | 0.35 | 0.77 | 0.01 |
| Unemployed | Percent unemployed | 1.0 | 1.12 | 0.92 | 1.37 | 0.34 |
| illiterate | Percent illiterate | 1.0 | 0.98 | 0.94 | 1.02 | 0.43 |

| Variable | Label | | LE(1) | | | |
|------------|--|-----|------------|------|------|--------|
| | | | Odds ratio | 1 | u | Prob |
| Diesel_PM | Diesel PM emissions from on-road and non-road sources | 1.0 | 0.99 | 0.97 | 1.01 | 0.53 |
| Fami11 | Two-parent family householders | 1.0 | 0.80 | 0.70 | 0.91 | 0.006 |
| Fami12 | single-parent family householders | 1.0 | 1.55 | 1.12 | 2.14 | 0.027 |
| Fami21 | Two-parent families | 1.0 | 1.32 | 1.16 | 1.51 | 0.0004 |
| Fami22 | single-parent families | 1.0 | 0.52 | 0.39 | 0.71 | 0.0004 |
| Fami32 | Single father or two-parent families | 1.0 | 0.83 | 0.75 | 0.92 | 0.0032 |
| Kinsured | percent of under age 18 insured | 1.0 | 0.84 | 0.74 | 0.96 | 0.03 |
| PM25 | annual mean PM 2.5 concentrations | 1.0 | 0.88 | 0.80 | 0.97 | 0.02 |
| Pesticides | Total pounds of selected active pesticide ingredients used in production-agriculture | 1.0 | 1.00 | 1.00 | 1.00 | 0.07 |
| Quart3 | college dormitories including college quarters off campus | 1.0 | 1.03 | 0.95 | 1.10 | 0.55 |
| bed7 | beds for acute care hospitals either present or absent | 1.0 | 2.11 | 1.08 | 4.12 | 0.07 |
| facip1 | specialty clinics either present or absent | 1.0 | 2.06 | 1.07 | 3.96 | 0.07 |
| facip3 | inpatient mental health either present or absent | 1.0 | 2.55 | 0.62 | 10.5 | 0.28 |
| farm | area is cultivated or not | 1.0 | 0.30 | 0.09 | 1.02 | 0.11 |
| gini | | 0.1 | 1.67 | 1.12 | 2.49 | 0.03 |
| openspace | area is open space or not | 1.0 | 2.55 | 1.39 | 4.66 | 0.01 |
| park | Percent of population within 1/2 mile of park, beach, open space, or coastline | 1.0 | 1.00 | 0.99 | 1.00 | 0.34 |
| race2 | Black | 1.0 | 0.95 | 0.90 | 1.00 | 0.11 |
| race4 | Asian and Pacific Islander | 1.0 | 1.03 | 1.01 | 1.06 | 0.03 |
| snap | percent of household receiving snap | 1.0 | 0.99 | 0.93 | 1.05 | 0.77 |

Table 8: Phase-II Results from "best" multiple logistic regression model controlling for phase-I variables (2010)

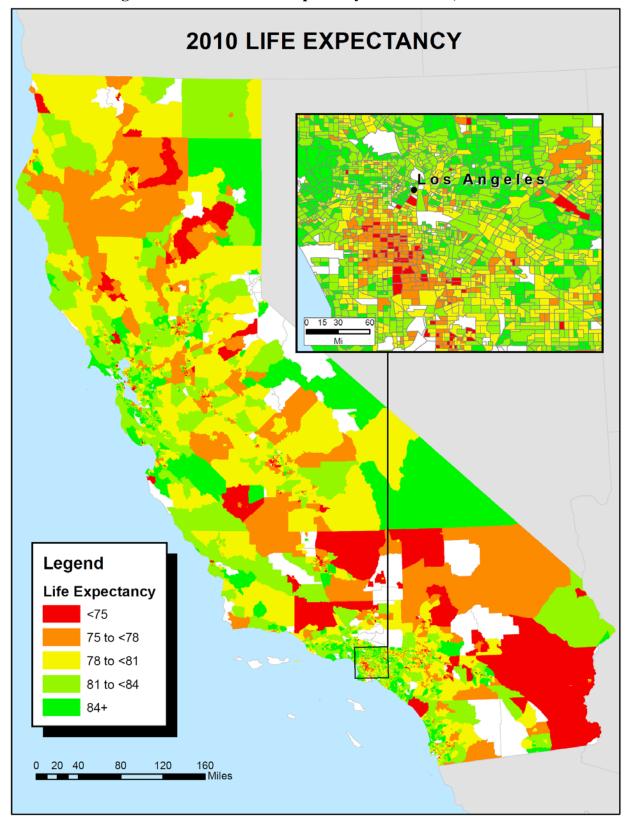


Figure 1: Census tract life expectancy in California, 2010

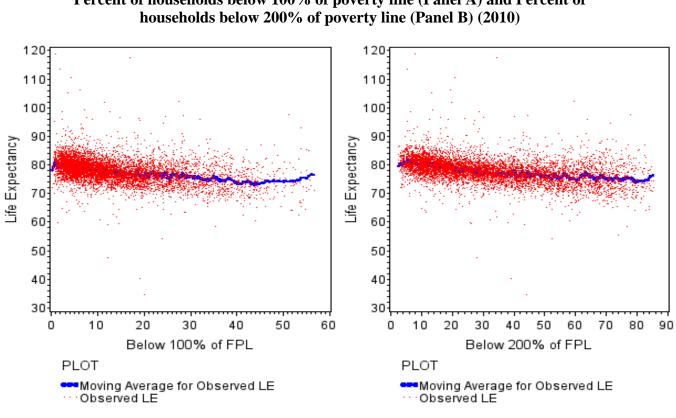
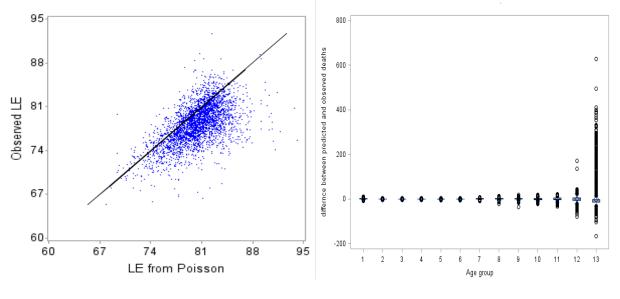


Figure 2: Life expectancy and percent poverty by census tract: Percent of households below 100% of poverty line (Panel A) and Percent of households below 200% of poverty line (Panel B) (2010)

Panel A

Panel B

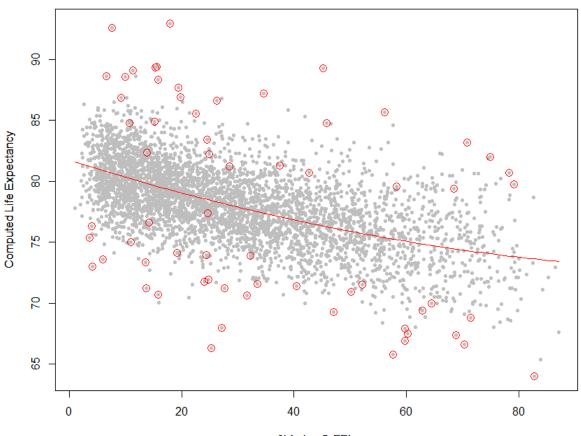
Figure 3: Observed life expectancy and life expectancy from Poisson modeling. Scatterplot of life expectancy (Panel A) and difference of observed and predicted life expectancy from Poisson Models (Panel B)



Panel A

Panel B

Figure 4: Association between life expectancy and percent below 2x poverty level for selected California census tracts meeting inclusion criteria. Census tracts surrounded by red circles were those considered "outliers" in the analysis.



All Populated Census Tracts with Population>=1250, Deaths>=20, 0<SE(LE)<= 2.5

% below 2xFPL Identify CTs having LE with |sresidual|>=2.576

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APPENDIX A: Variable lists

| Phase | Variable | Data Source | Geographic Scale | Year |
|-------|---|---|------------------|--|
| 1 | Deaths | California Department of Public Health | Census tract | 1999-2001 |
| 1 | % of population aged >= 25 years by educational attainment | American Community Survey | Census tract | City (1, 3 years); Census tract (5 years) |
| 1 | Gender | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2000 |
| 1 | Age Group Distribution | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2000 |
| 1 | Income-to-Poverty Ratio | <u>U.S. Census Bureau, American Community</u> <u>Survey: American FactFinder: st00007.uf3</u> | Census tract | 2000 |
| 1 | Marital status | U.S. Census Bureau, American Community Survey: American FactFinder: Table DP-2 | Census tract | 2000 |
| 1 | Immigration status | American Community Survey | Census tract | 2000 |
| 1 | Hispanic Origin | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2000 |
| 2 | Race | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2000 |
| 2 | Percent of unhealthy days | <u>http://www.epa.gov/airdata/ad_rep_aqi.ht</u> <u>ml</u> | County | 2000 |
| 2 | % of children less than 18 reported with neglect, physical, or sexual abuse | <u>Child Welfare Dynamic Report System .</u> <u>University of California, Berkeley and</u> <u>Dept. of Social Services.</u> | Census Tract | 2000 |
| 2 | Group Quarters | American Community Survey | Census Tract | 2000 |
| 2 | Cancer risk due to air pollutants (per million) | National Air Toxic Assessment (http://www.epa.gov/ttn/atw/nata1999/tabl es.html) | Census Tract | 1999 |
| 2 | Respiratory risk due to air pollutants (HI) | National Air Toxic Assessment (http://www.epa.gov/ttn/atw/nata1999/tabl es.html) | Census Tract | 1999 |
| 2 | Neurological risk due to air pollutants (HI) | National Air Toxic Assessment | Census Tract | 1999 |
| 2 | Time traveled to work | American Community Survey | Census Tract | 2000 |
| 2 | Rural vs Urban Location | American Community Survey | Census Tract | 2000 |
| 2 | percent of residents that walk or bike to work | American Community Survey. Detailed data sets, annual, 3-year, 5-year | Census tract | 2000 |
| 2 | Access to healthy food | <u>MATCH 2011</u> | County | 2008 |
| 2 | Retail food environment index (e.g. number of fast- food restaurants and convenience stores/total number of supermarkets and produce vendors | <u>California Board of Equalization</u> | Census tract | 2011 |
| 2 | insured Population | American Community Survey | County | 2000 |
| 2 | Insured children under 20 | California Healthy Kids Survey | County | 2001 |
| 2 | Immunization status | California Healthy Kids Survey | County | 2001 |

Table A1: Data Sources (2000)

| 2 | 3 <i>x</i> - 1 - 1 11 - 1 - 1-1 | NATOR 2011 | <u> </u> | 2000 |
|---|---|--|--------------|-----------|
| 2 | Mental public health provider ratio | <u>MATCH 2011</u> | County | 2008 |
| 2 | Primary care providers rate | <u>MATCH 2011</u> | County | 2006 |
| 2 | household type | American Community Survey. Detailed data sets, annual, 3-year, 5-year and California Department of Housing and Community Development. Building Blocks for Effective Housing Elements, housing needs | Census tract | 2000 |
| 2 | Percent of Household income spent on rent or mortgage using benchmarks of >30% (burdened) and >50% severely burdened | American Community Survey. Detailed data sets, annual, 3-year, 5-year and California Department of Housing and Community Development. Building Blocks for Effective Housing Elements. Housing needs | Census tract | 2000 |
| 2 | Income Inequality: Gini coefficient describing the amount of total annual community income generated by the number of households | American Community Survey. Detailed data sets, annual, 3-year, 5-year. U.S Bureau of Census. | Census tract | 2000 |
| 2 | employment status | American Community Survey. Detailed data sets, annual, 3-year, 5-year. U.S Bureau of Census. AND 29. Monthly Labor Force Data for Cities and Census Designated Places (CDP). Employment Development Department. | Census tract | 2000 |
| 2 | % of households in overcrowded (>= 1.01 persons/room) and severely overcrowded (>= 1.5 persons per room) conditions | American Community Survey. Detailed data sets, annual, 3-year, 5-year: AND California Department of Housing and Community Development. Building Blocks for Effective Housing Elements. Housing needs | Census tract | 2000 |
| 2 | recreation facility rate | <u>MATCH 2011</u> | County | 2008 |
| 2 | Sex ratio at birth | California Healthy Kids Survey | County | 2000 |
| 3 | Mentally unhealthy days | <u>MATCH 2011</u> | County | 2003-2009 |
| 3 | Diabetic | <u>MATCH 2011</u> | County | 2003-2006 |
| 3 | Adults obesity | <u>MATCH 2011</u> | County | 2006-2008 |
| 3 | Smokers | <u>MATCH 2011</u> | County | 2002-2008 |
| 3 | physical unhealthy days | <u>MATCH 2011</u> | County | 2002-2008 |
| 3 | physical inactivity | <u>MATCH 2011</u> | County | |
| 3 | Teen birth rate | <u>MATCH 2011</u> | County | 2000-2006 |
| 3 | Percent of population located <1/2 mile of a regional bus/rail/ferry and <1/4 mile local bus /light rail | University of California Davis (ULTRANS), University of California Berkeley (SafeTREC). Transit asset inventories (Link 1 and link 2). | Census Tract | Annual |

| 3 | % of | American Community Survey. Detailed | Census tract | Annual |
|---|--------------------------|--|--------------|-----------|
| | households/populations | data sets, annual, 3-year, 5-year. U.S | | |
| | near busy roadways | Bureau of Census. Federal Highway | | |
| | | Administration. FHWA Functional | | |
| | | Classification Guidelines. Department of | | |
| | | Transportation; 2008 in conjunction with | | |
| | | public or commercial GIS files of | | |
| | | roadways. | | |
| 3 | Hospitalization rate for | MATCH 2011 | County | 2005-2006 |
| | ambulatory care | | 2 | |
| 3 | HIV rate | MATCH 2011 | County | |
| _ | | | - | |
| 3 | STI rate | MATCH 2011 | County | |

Table A2: Data Sources (2010)

| Phase | variable | Data Source | Geographic Scale | Years/Frequency of Update |
|-------|---|--|----------------------|------------------------------|
| 1 | Deaths | California Department of Public Health | Census tract | • |
| 1 | % of population aged >= 25 years by educational attainment | American Community Survey. Detailed data sets, annual, 3-year, 5-year. U.S Bureau of Census, and Condition of education. U.S. Department of Education National Center for Educational Statistics. | Census tract | 2006-2010 |
| 1 | Gender | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2006-2010 |
| 1 | Age Group Distribution | U.S. Census Bureau. American Community Survey: American FactFinder Table DP-1 | Census tract | 2006-2010 |
| 1 | Income-to-Poverty Ratio | U.S. Census Bureau, American Community Survey: American FactFinder: st00007.uf3 | Census tract | 2006-2010 |
| 1 | Marital status | U.S. Census Bureau, American Community Survey: American FactFinder: Table DP-2 | Census tract | 2006-2010 |
| 1 | immigrant status | American Community Survey | Census tract | 2006-2010 |
| 1 | Hispanic Origin | <u>U.S. Census Bureau, American Community</u> Survey: American FactFinder Table DP-1 | Census tract | 2006-2010 |
| 2 | Race | U.S. Census Bureau, American Community Survey: American FactFinder Table DP-1 | Census tract | 2006-2010 |
| 2 | percent of tree canopy coverage (urban areas) | National Land Cover Database. USGS, EPA, NOAA, DOI, NASA, USFS, US Park Service (Multi-Resource Land Consortium) and other public and commercial satellite imaging projects. | Census tract | 2010 |
| 2 | Percent of unhealthy days | http://www.epa.gov/airdata/ad_rep_aqi.html | County/ Air basin | 2010 |
| 2 | % of children less than 18 reported with neglect, physical, or sexual abuse | <u>Child Welfare Dynamic Report System . University of</u> <u>California, Berkeley and Dept. of Social Services.</u> | Census Tract | 2010 |
| 2 | Group Quarters | American Community Survey | Census Tract | 2010 |
| 2 | Cancer risk due to air pollutants (per million) | National Air Toxic Assessment (http://www.epa.gov/ttn/atw/nata1999/tables.html) | Census Tract | 2005 |
| 2 | Respiratory risk due to air pollutants (HQ) | <u>National Air Toxic Assessment</u> (http://www.epa.gov/ttn/atw/nata1999/tables.html) | Census Tract | 2005 |
| 2 | Neurological risk due to air pollutants (HQ) | National Air Toxic Assessment | Census Tract | 2005 |
| 2 | Time traveled to work | American Community Survey | Census Tract | 2006-2010 |
| 2 | Rural vs Urban Location | American Community Survey | Census Tract | 2006-2010 |
| 2 | percent of residents that walk or bike to work | <u>American Community Survey. Detailed data sets.</u> annual, 3-year, 5-year | Census tract | 2006-2010 |
| 2 | Access to healthy food | MATCH 2013 | county | 2006 |
| 2 | Retail food environment index (eg number of fast-food restaurants and convenience stores/total number of supermarkets and produce vendors | <u>California Board of Equalization</u> | Census tract | 2011 |
| 2 | Uninsured Population | American Community Survey | County | 2010 |
| 2 | Insured children under 20 | California Healthy Kids Survey | County | 2009 |
| 2 | Immunization status | <u>California Healthy Kids Survey</u> | County | 2010 |
| 2 | Mental public health provider ratio | <u>MATCH 2013</u> | county | 2011-2012 |
| 2 | Primary care providers rate | <u>MATCH 2013</u> | county | 2006 |
| 2 | household type | <u>American Community Survey. Detailed data sets.</u> <u>annual, 3-year, 5-year and California Department of</u> <u>Housing and Community Development. Building</u> | Census tract | 2006-2010 |
| 2 | Percent of Household | Blocks for Effective Housing Elements, housing needs American Community Survey. Detailed data sets, | Census tract | 2006-2010 |

| 3 | Percent of population located <1/2 mile of a regional bus/rail/ferry | University of California Davis (ULTRANS), University of California Berkeley (SafeTREC). Transit asset inventories (Link 1 and link 2). | Census Tract | Annual |
|---|---|---|--------------|-----------|
| 3 | Teen birth rate | MATCH 2013 | County | 2004-2010 |
| 3 | physical inactivity | MATCH 2013 | County | 2009 |
| 3 | physical unhealthy days | MATCH 2013 | County | 2005-2011 |
| 3 | Smokers | MATCH 2013 | county | 2005-2011 |
| 3 | Adults obesity | <u>MATCH 2013</u> | county | 2009 |
| 3 | Diabetic | <u>MATCH 2013</u> | county | 2010 |
| 3 | Mentally unhealthy days | MATCH 2013 | county | 2005-2011 |
| 2 | % population within 1/2 mile of park, beach or open space | CalEnviroScreen | ZCTA | 2009-2010 |
| 2 | Pollution Burden | <u>CalEnviroScreen</u> | ZCTA | 2009-2010 |
| 2 | Traffic Density | <u>CalEnviroScreen</u> | ZCTA | 2009-2010 |
| 2 | Pesticides | <u>CalEnviroScreen</u> | ZCTA | 2009-2010 |
| 2 | Diesel PM | CalEnviroScreen | ZCTA | 2009-2010 |
| 2 | PM 2.5 (ZCTA) | <u>CalEnviroScreen</u> | ZCTA | 2009-2010 |
| 2 | psychiatric services Number of nursing homes | Development: Healthcare Atlas California Office of Statewide Health Planning and Development: Healthcare Atlas | Census Tract | 2012 |
| 2 | psychiatric services Number of out-patient | Development: Healthcare Atlas California Office of Statewide Health Planning and | Census Tract | 2012 |
| 2 | Number of in-patient | California Office of Statewide Health Planning and | Census Tract | 2012 |
| 2 | services Number of acute care hospitals | California Office of Statewide Health Planning and Development: Healthcare Atlas | Census Tract | 2012 |
| 2 | clinics number of hospice/home care | Development: Healthcare Atlas California Office of Statewide Health Planning and Development: Healthcare Atlas | Census Tract | 2012 |
| 2 | number of specialty | California Office of Statewide Health Planning and | Census Tract | 2012 |
| 2 | Number of free/community clinics | <u>California Office of Statewide Health Planning and</u> Development: Healthcare Atlas | Census Tract | 2012 |
| 2 | to cost PM 25 | - | county | 2010 |
| 2 | social support could not see doctor due to cost | <u>MATCH 2013</u> | county | 2005-2011 |
| 2 | without emotional or | <u>MATCH 2013</u> | county | 2005-2010 |
| 2 | child care cost | California Healthy Kids Survey | county | 2010 |
| 2 | child blood lead level | California Healthy Kids Survey | county | 2010 |
| 2 | Sex ratio at birth | California Healthy Kids Survey | County | 2010 |
| 2 | persons/room) and severely overcrowded (>= 1.5 persons per room) conditions recreation facility rate | Housing and Community Development. Building Blocks for Effective Housing Elements. Housing needs | County | 2010 |
| 2 | % of households in overcrowded (>= 1.01 | annual, 3-year, 5-year, U.S Bureau of Census. American Community Survey. Detailed data sets. annual, 3-year, 5-year: AND California Department of | Census tract | 2006-2010 |
| 2 | coefficient describing the amount of total annual community income generated by the number of households employment status | annual, 3-year, 5-year. U.S Bureau of Census. American Community Survey. Detailed data sets, | census tract | 2006-2010 |
| 2 | benchmarks of >30% (burdened) and >50% severely burdened Income Inequality: Gini | Blocks for Effective Housing Elements. Housing needs American Community Survey. Detailed data sets. | Census tract | 2006-2010 |
| | income spent on rent or mortgage using | annual, 3-year, 5-year and California Department of Housing and Community Development, Building | | |

| | and <1/4 mile local bus /light rail | | | |
|---|--|---|--------------|-----------|
| 3 | % of households/populations near busy roadways | American Community Survey. Detailed data sets, annual, 3-year, 5-year. U.S Bureau of Census. Federal Highway Administration. FHWA Functional Classification Guidelines. Department of Transportation; 2008 in conjunction with public or commercial GIS files of roadways. | Census tract | Annual |
| 3 | Hospitalization rate for ambulatory care | MATCH 2013 | county | 2005-2006 |
| 3 | HbA1c | MATCH 2013 | County | 2009 |
| 3 | STI rate | MATCH 2013 | County | 2010 |
| 2 | SNAP | American Community Survey | Census tract | 2007-2011 |
| 2 | School Enrollment | American Community Survey | Census tract | 2007-2011 |
| 2 | Transportation to work | California Department of Public Health | Census tract | 2006-2010 |